

The Impact of Investor Sentiment on Energy and Stock Markets-Evidence : China and Hong Kong

Liang-Chun Ho*

Received: January 14, 2014. Revised: February 09, 2014. Accepted: March 17, 2014.

Abstract

Purpose - The oil price affects company value, which is the present value of the expected cash flow, by affecting the discount rate and cash flow. This study examines the nonlinear relationships between oil price and stock price using the AlphaShares Chinese Volatility Index as the threshold.

Research design, data, and methodology - Data comprise daily closing values of the Shanghai Stock Exchange Composite Index, Shenzhen Stock Exchange Composite Index, and Hang Seng Index of ChinaWest Texas Intermediate crude oil spot price and AlphaShares Chinese Volatility Index from May 25, 2007 to May 24, 2012. The Threshold Error Correction Model is used.

Results - The results demonstrate different relationships between the stock price index and oil price under different investor sentiments; however, the stock price index and oil price could adjust to a long-term equilibrium the long-term causality tests between them were all significant.

Conclusions - The relationship between the WTI and HANG SENG Index is more significant than the Shanghai Composites Index and Shenzhen Composite Index, when using the AlphaShares Chinese Volatility Index (ASC-VIX) as the investor sentiment variable and threshold.

Keywords: Oil Price, Stock Price, Investor Sentiment, TECM, China, Hong Kong.

JEL Classifications: G00, G10, F00.

1. Introduction

Oil price affects economy, such as output, employment, inflation, interest rate, consumption, economic growth and re-

cession, since oil is a major production factor (Hamilton, 1983; Bernake, 1983; Gisser and Goodwin, 1986; Chen, Roll, and Ross, 1986; Pindyck, 1991; Rotemberg and Woodford 1996; Jones and Kaul, 1996; Sadorsky, 1999; Park and Ratti, 2008, Apergis and Miller, 2009; El Dedi Aroui, Jouini, and Nguyen, 2011). However, the impact of oil shock on economy downturn is still inclusive (Hooker, 1966). Literature shows the nonlinear relationships between oil prices and the economy (Lee et al., 1996; Hamilton, 1996; Hungtinton, 1998). The increase of oil price has bigger impact on economy downturn than the decrease of oil price on economy growth.

Oil price affects company value since the company value is the present value of the expected cash flow. Therefore, oil price affects stock price by two factors, discount rate and cash flow. Inflation rate will increase if oil price increases. This will lead to higher interest rate and discount rate. Stock price has negative relationships with discount rate. However, investors receive two kinds of returns from stock investment, dividend and capital gain. A positive expected dividend gain plus expected capital gain tends to have the stock price increase. Therefore, high inflation may not lead to lower stock price. During high inflation period, return from saving account could be negative. It is possible that the stock price increases during high oil price and high inflation period if the expected capital gain is high. Some empirical studies show the negative relationship between oil price and stock price (Kling, 1985; Jones and Kaul, 1996; Faff and Brailsford, 1999; Papapetrou, 2001; Sadorsky, 2001; Driesprong, Jacobsen and Maat, 2008). However, some literature shows weak or no relationship between oil price and stock price (Huang et al, 1996; El-Sharif et al, 2005; Cong, Wei, and Liao, 2008). Oil price shocks have negative impacts on the oil import countries, and have positive impacts on the oil export countries (Park and Ratti, 2008). Oil price increase has positive impact on the oil and gas company stocks (Sadorsky, 2001).

The conditional expected excess market return and the conditional variance of the market are related based on the intertemporal asset pricing model (Merton, 1973). However, the conditional variance is unobservable (Ghysels et al., 2005). This makes it hard to have a unified conclusion about the relationship between the conditional expected excess return and conditional variance. Literature show positive relationships between

* Corresponding Author, Assistant Professor, Department of Applied Financial, Hsiuping University of Science and Technology, Taiwan.
E-Mail: konah@mail.hust.edu.tw.

conditional expected excess return and conditional variance (French et al, 1987; Ghysels et al, 2005; Guo and Whitelaw, 2006; Guo and Neely, 2008 and Lundblad, 2007) and negative and weak relationships between conditional expected excess return and conditional variance (Baillie and DeGennaro, 1990; Campbell and Hentschel, 1992; Brandt and Kang, 2004; Goyal and Santa-Clara, 2003; Lettau and Ludvigson, 2003; and Li et al. 2005). VIX index, which contains information about future excess market return and variance, can be used to estimate the conditional variance (Day and Lewis, 1992; Blair et al., 2001; Guo and Whitelaw, 2006; Banerjee et al., 2007; Kanas, 2012).

The Chicago Board Options Exchange Volatility Index (VIX) is an indicator of future 30 days U.S. stock market volatility. It is considered to be an investor fear gauge or investor sentiment in the U.S. stock markets. Literature shows that VIX is called the investor fear gauge. High levels of VIX coincide with high degrees of market turmoil in the U.S. (Whaley, 2000). Some literature shows high level of VIX index followed by positive future stock market return and low level of VIX index followed by negative future stock market return (French et al., 1987; Fleming et al., 1995; Giot, 2005; Guo and Whitelaw, 2006). This indicates that oversold market tends to have high level VIX index. Oversold stock market leads to a positive relationship between VIX index and future stock market return (Giot, 2005). Some literature shows negative contemporaneous relationship between VIX changes and stock market return (Fleming et al., 1995). This indicates that VIX index and stock market index have inverse relationship. Since VIX index is mean reverting, a negative contemporaneous relationship between VIX index changes and stock market returns may be followed by a positive relationship between VIX index changes and past stock market returns (Guo and Wohar, 2006). Some literature shows that the VIX index and stock market return have asymmetric relationship. An increase of VIX index from a negative stock market return is bigger than the decrease of VIX index from a similar positive stock market return (Schwert, 1990; Flemming et al., 1995). VIX index has significantly suppressing effect on oil prices in the long run (Sari, Soytaş and Hacihasanoglu, 2011). It will be interesting to see if VIX can also be used as an investor sentiment in Chinese market. AlphaShares Chinese Volatility Index (ASC-VIX) is used in this paper. In addition, from our knowledge, there is no paper that uses the VIX index as the threshold to study the nonlinear relationships between oil price and stock price. This paper studies the nonlinear relationships between oil price and stock price by using the AlphaShares Chinese Volatility Index (ASC-VIX) as the threshold.

This paper is organized as follows. Section 2 presents the data used in the study. Section 3 briefly describes the empirical methodology. Section 4 shows the empirical results and section 5 concludes this study.

2. Data

Daily closing values of Shanghai Stock Exchange Composite

Index (SSE Composite Index), Shenzhen Stock Exchange Composite Index (SZSE Composite Index) and Hang Seng Index (HSI) of China, West Texas Intermediate (WTI) crude oil spot price, and AlphaShares Chinese Volatility Index (ASC-VIX) are used in this study. The research period is from May 25, 2007 to May 24, 2012. And data of the variables is taken from Taiwan Economic Journal (TEJ), Energy Information Administration (EIA) and Bloomberg¹.

AlphaShares China Volatility Index is a measurement of the implied volatility of options on major China equity indexes. Chicago Board Option Exchange Volatility Index (VIX) is the market volatility expectation. AlphaShare Chinese Volatility Index is a measurement of the implied volatility of options on major China equity indexes, including Hang Seng Index and FTSE/Xinhua 25 index. It is calculated by a formula similar to the Chicago Board Options Exchange Volatility Index (VIX). AlphaShares China Volatility Index can be used to represent the market volatility expectation of Chinese market.

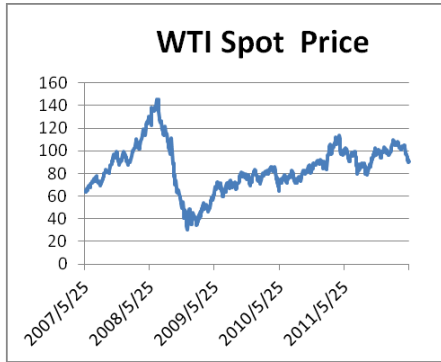
Figures 1 to 8 show the historical chart for each variable. Table 1 presents some descriptive statistics, such as mean, standard deviation, maximum, minimum, skewness, kurtosis, and Jarque-Bera value. Among three stock indices and WTI, WTI has the smallest standard deviation and HANG SENG Index has the largest standard deviation. About skewness statistics, only SHANGHAI COMPOSITE Index is right-tailed. And the Jarque-Bera tests show that all variables reject the null hypothesis of normal distribution at 1% significance level.

<Table 1> descriptive statistics of Variables

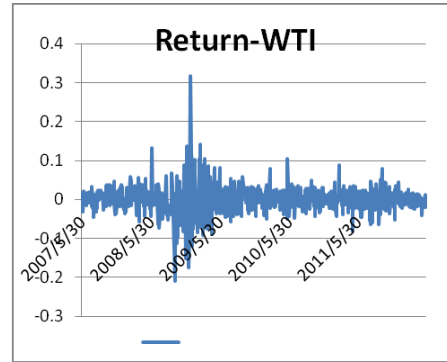
	HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index	WTI
Mean	20790.240	2950.098	1067.540	83.129
Median	21099.770	2797.547	1095.158	82.085
Maximum	31638.220	6036.281	1570.396	145.310
Minimum	11344.580	1719.774	466.052	30.280
Std. Dev.	3220.191	779.880	212.422	19.290
Skewness	-0.534	1.727	-0.553	-0.025
Kurtosis	4.090	6.035	3.207	3.414
Jarque-Bera	98.185***	881.380***	52.825***	7.253***

1. ***, **, and * denote the significant levels at 1%, 5%, and 10%.

1) <http://www.bloomberg.co.jp/apps/quote? T=jp09/quote.wm&ticker = ASCNCHIX: IND>



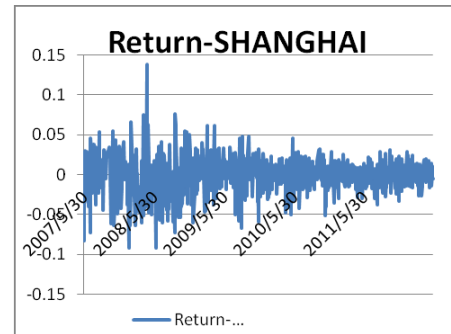
<Figure 1> 2007/5/25-2012/5/24 West Texas Intermediate crude oil



<Figure 2> Return of West Texas Intermediate crude oil



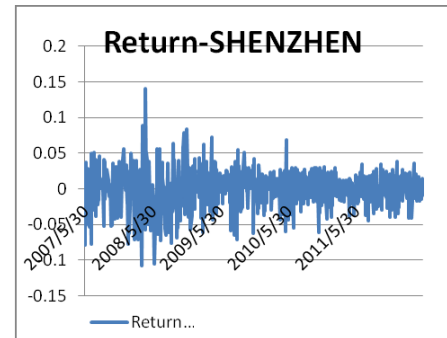
<Figure 3> 2007/5/25-2012/5/24 Shanghai Composites Index



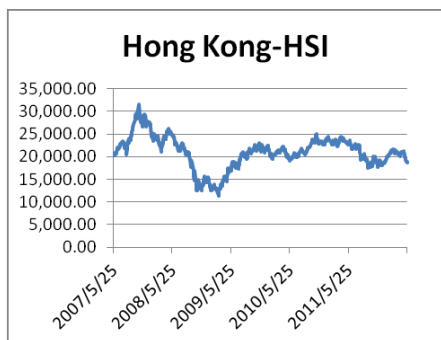
<Figure 4> Return of Shanghai Composites Index



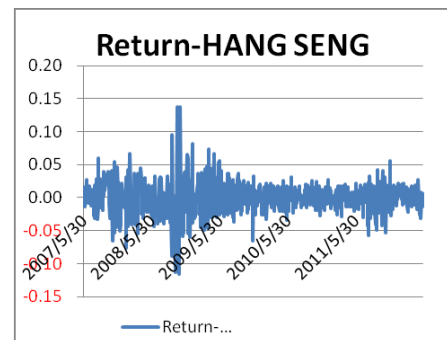
<Figure 5> 2007/5/25-2012/5/24 Shenzhen Composite Index



<Figure 6> Return of Shenzhen Composite Index



<Figure 7> 2007/5/25-2012/5/24 HANG SENG Index



<Figure 8> Return of HANG SENG Index

3. Research Method

3.1. Nonlinear Unit Root Test: Kapetanios, Shin, and Snell (KSS) Test

The Augmented Dickey-Fuller test (Sims, 1988) and Phillips-Perron test (Phillips and Perron, 1988) are used for the linear unit root test. The ADF and PP tests are for the null hypothesis that a time series is $I(1)$. The Kwiatkowski, Phillips, Schmidt and Shin (1992) KPSS test is used to reinforce the results from ADF and PP tests. The KPSS stationary test is for the null hypothesis that a time series is $I(0)$. The Kapetanios, Shin, and Snell (2003) KSS nonlinear unit root test is used for the nonlinear stationary test.

The KSS test is to detect the presence of non-stationary against a nonlinear but globally stationary exponential smooth transition autoregressive (ESTAR) process. The model is expressed as follows.

$$\Delta X_t = \alpha X_{t-1} \{1 - \exp(-\gamma X_{t-1}^2)\} + \varepsilon_t \quad (1)$$

where X_t is the time series data. ε_t is an independently identically distributed, iid, error term with zero mean and constant variance. And $\gamma \geq 0$ is the transition parameter of the ESTAR model and governs the speed of transition. Under the null hypothesis, X_t follows a linear unit root process. And X_t follows a nonlinear stationary ESTAR process under the alternative hypothesis. One shortcoming in this framework is that the parameter γ is not identified under the null hypothesis. Replacing $\{1 - \exp(-\gamma X_{t-1}^2)\}$ in equation (1) by a first-order Taylor series expansion around $\gamma = 0$, the following auxiliary model can be derived (Kapetanios, Shin, and Snell, 2003).

$$\Delta X_t = c + \theta X_{t-1}^3 + \sum_{i=1}^k \beta_i \Delta X_{t-i} + v_t, \quad t = 1, 2, \dots, T \quad (2)$$

$$\begin{cases} H_0: \theta = 0 \\ H_1: \theta < 0 \end{cases} \quad (3)$$

Under this framework, the null hypothesis and the alternative hypothesis are expressed as $\theta = 0$ (non-stationary) against $\theta < 0$ (nonlinear ESTAR stationary). If it is rejected then it is nonlinear stationary. If it cannot be rejected then it is non-stationary.

3.2. Threshold Autoregressive / Momentum-Threshold Autoregressive cointegration tests

Nonlinear asymmetric adjustment processes may exist between the nonstationary variables in the linear models. Nonlinear models are better solutions to capture nonlinear cointegration relationship between variables. (Enders and Granger, 1998; Enders and Siklos, 2001). Threshold Autoregressive (TAR)/Momentum-Threshold Autoregressive (MTAR) cointegration tests are used to study nonlinear cointegration relationships between

the variables.

Unit root test can be used to test those variables in the vector $\{X_{1t}, X_{kt}\}$ are integrated of order 1 (Engle and Granger, 1987).

$$X_t = c + \beta_t X_{t-1} + \varepsilon_t \quad (4)$$

$$\Delta \varepsilon_t = \rho \varepsilon_{t-1} + v_t \quad (5)$$

Cointegration implies that ε_t is stationary with zero mean and $\rho = 0$. The change in ε_t equals ρ multiplied by ε_{t-1} regardless of whether ε_{t-1} is positive or negative (Johansen, 1995).

$$\Delta X_t = \eta X_{t-1} + \varepsilon_t \quad (6)$$

Where X_t is a $(k \times 1)$ vector, η is a $(k \times k)$ matrix, and ε_t is a $(k \times 1)$ vector of normally distributed disturbances that may be contemporaneously correlated. The Johansen procedure is to test the null hypothesis that the rank of η equals zero. Under the alternative hypothesis, $\text{rank}(\eta) \neq 0$, the adjustment process is symmetric around $X_t = 0$ such that for any $X_t \neq 0$, ΔX_{t+1} always equals ηX_t .

The implicit assumption of symmetric adjustment is problematic if the adjustment towards the long-run equilibrium relationship is not linear. Enders and Granger (1998), and Enders and Siklos (2001) introduce asymmetric adjustment by letting the deviations from the long-run equilibrium in equation (4) behave as a Threshold Autoregressive (TAR) process:

$$\Delta \varepsilon_t = \rho_1 \varepsilon_{t-1} I_t + \rho_2 \varepsilon_{t-1} (1 - I_t) + v_t \quad (7)$$

$$I_t = \begin{cases} 1 & , \varepsilon_{t-1} > \tau, \\ 0 & , \varepsilon_{t-1} \leq \tau, \end{cases} \quad (8)$$

$$\begin{cases} H_0: \rho_1 = \rho_2 = 0 \\ H_1: \rho_1 \neq \rho_2 \neq 0, \end{cases} \quad \begin{cases} H_0: \rho_1 \neq \rho_2 \\ H_1: \rho_1 = \rho_2 \end{cases} \quad (9)$$

Asymmetric adjustment is implied by different values of ρ_1 and ρ_2 when ε_{t-1} is positive, the adjustment is $\rho_1 \varepsilon_{t-1}$, and if ε_{t-1} is negative, the adjustment is $\rho_2 \varepsilon_{t-1}$. A sufficient condition for stationary of $\{\varepsilon_t\}$ is $-2 < (\rho_1, \rho_2) < 0$. If the $\{\varepsilon_t\}$ sequence is stationary, the least squares estimates of ρ_1 and ρ_2 have an asymptotic multivariate normal distribution. The adjustment is symmetric, $\rho_1 = \rho_2$, if the null hypothesis $\rho_1 = \rho_2 = 0$ is rejected. Then the standard F-test is more powerful in the case of symmetric adjustment.

If the adjustment is symmetric, $\rho_1 = \rho_2$, the Engle-Granger (1987) test for cointegration is a special case of equation (7). The exact nature of the non-linearity may not be known, it is also possible to allow the adjustment to depend on the change of ε_{t-1} instead of the level of ε_{t-1} . The momentum threshold autoregressive (MTAR) model can be used to test the null hypothesis of a unit root (Enders and Granger, 1998; Enders and Siklos, 2001)

$$\Delta \varepsilon_t = \rho_1 \varepsilon_{t-1} I_t + \rho_2 \varepsilon_{t-1} (1 - I_t) + \sum \beta_i \Delta \varepsilon_{t-i} + v_t \quad (10)$$

$$I_t = \begin{cases} 1, \Delta \varepsilon_{t-1} > \tau \\ 0, \Delta \varepsilon_{t-1} \leq \tau \end{cases} \quad (11)$$

$$\begin{cases} H_0: \rho_1 = \rho_2 = 0 \\ H_1: \rho_1 \neq \rho_2 \neq 0, \end{cases} \quad \begin{cases} H_0: \rho_1 \neq \rho_2 \\ H_1: \rho_1 = \rho_2 \end{cases} \quad (12)$$

3.3. Threshold Error Correction Model

If there are some relationships between energy markets and stock market, the model is as follows:

$$Y = \alpha + \beta X_i + \varepsilon \quad (13)$$

where Y is stock price, X is WTI spot price.

Whether AlphaShares Chinese Volatility Index (ASC-VIX) can be used as an investor sentiment in Chinese market is tested in this paper. ASC-VIX index contains information about future excess market return and variance and can be used to estimate the conditional variance. The ASC-VIX is used as the threshold in this paper to study the nonlinear relationships between oil price and stock price. Therefore equation (13) can be revised to equations (14) and (15) as follows:

$$Y = \alpha_0 + \alpha_1 X + \varepsilon, \quad \text{if } ASC - VIX > \gamma \quad (14)$$

$$Y = \beta_0 + \beta_1 X + \varepsilon, \quad \text{if } ASC - VIX \leq \gamma \quad (15)$$

Following Enders and Granger (1998), Enders and Siklos (2001), the threshold error correction model (TECM), the equation (14) and (15) was changed to (16) and (17), to research the nonlinear relationships between stock price and oil price.

$$\begin{aligned} \Delta STOCK_i = & (\alpha_0 + \alpha_1 \Delta STOCK_{i,t-1} + \alpha_2 \Delta STOCK_{i,t-2} + \alpha_3 \Delta WTI_{i,t-1} \\ & + \alpha_4 \Delta WTI_{i,t-2} + \alpha_5 ECT_{i,t-1}) \times I \\ & + (\beta_0 + \beta_1 \Delta STOCK_{i,t-1} + \beta_2 \Delta STOCK_{i,t-2} + \beta_3 \Delta WTI_{i,t-1} \\ & + \beta_4 \Delta WTI_{i,t-2} + \beta_5 ECT_{i,t-1}) \times (1 - I) + \varepsilon_i \end{aligned} \quad (16)$$

$$\begin{aligned} \Delta WTI_i = & (\alpha_0 + \alpha_1 \Delta STOCK_{i,t-1} + \alpha_2 \Delta STOCK_{i,t-2} + \alpha_3 \Delta WTI_{i,t-1} \\ & + \alpha_4 \Delta WTI_{i,t-2} + \alpha_5 ECT_{i,t-1}) \times I \\ & + (\beta_0 + \beta_1 \Delta STOCK_{i,t-1} + \beta_2 \Delta STOCK_{i,t-2} + \beta_3 \Delta WTI_{i,t-1} \\ & + \beta_4 \Delta WTI_{i,t-2} + \beta_5 ECT_{i,t-1}) \times (1 - I) + \varepsilon \end{aligned} \quad (17)$$

$$I = \begin{cases} 1, & ASC - VIX > \gamma, \\ 0, & ASC - VIX \leq \gamma, \end{cases} \quad (18)$$

Where ΔWTI and $\Delta STOCK_i$ are the percentage change of the oil price and the stock indexes, ASC-VIX is AlphaShares China Volatility Index. TECM model is used to study the asymmetric relationship between oil prices and stock prices index under different AlphaShares China Volatility Index conditions. The process of interaction between oil prices and stock prices can be analyzed. The short-term causality and long-term causality will be studied.

4. Empirical Studies

4.1. Unit Root Test

The results of the Augmented Dickey-Fuller (ADF) tests, Phillips and Perron (PP) tests, and Kwiatkowski, Phillips, Schmidt and Shin (KPSS) test are shown in Table 2. The results show that HANG SENG Index, Shanghai Composites Index, Shenzhen Composite Index and WTI oil price are all non-stationary in levels but become stationary in the first differences. Table 3 presents the KSS nonlinear stationary test results. Table 3 indicates that stock price indexes of China and Hong Kong and WTI oil price series are no stationary, and become stationary in the first difference.

The results imply that stock price indexes of China and Hong Kong and WTI oil price are all non-stationary in levels but become stationary in the first differences. Stock price indexes and dividends are integrated of order one, I (1). These results indicate that all the oil price and stock indexes exhibit the unit root. The nonlinear cointegration test is used to examine the long-term equilibrium relationships.

<Table 2> Linear Unit Root Test

Panel A: Level		HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index	WTI
Intercept	ADF	0.278	0.392	0.344	0.285
	PP	0.254	0.500	0.268	0.244
	KPSS	0.271	1.214***	0.349**	0.893***
Intercept And Trend	ADF	0.593	0.705	0.661	0.567
	PP	0.562	0.635	0.570	0.503
	KPSS	0.257***	0.267***	0.293***	0.258***
Panel B: 1st difference		HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index	WTI
Intercept	ADF	0.000***	0.000***	0.000***	0.000***
	PP	0.000***	0.000***	0.000***	0.000***
	KPSS	0.067	0.094	0.077	0.060
Intercept And Trend	ADF	0.000***	0.000***	0.000***	0.000***
	PP	0.000***	0.000***	0.000***	0.000***
	KPSS	0.063	0.070	0.077	0.061

1. ADF and PP: p-values, KPSS: LM-Stat.

2. ***: denotes rejection of the Null Hypothesis (H0) at the 1% level

<Table 3> Nonlinear Unit Root Test - KSS Test

	HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index	WTI
T-statistic	-1.923582 (-2.82)	-1.645503 (-2.82)	-1.798411 (-2.82)	-1.994921 (-2.82)
lag	1	1	1	1
Case	1	1	1	1

1. ***: denotes rejection of the Null Hypothesis (H0) at the 1% level.

2. (): Critical values are taken from Kapetanios, Shin, and Snell (2003) Table1.

3. Case 1, Case 2 and Case 3 refer to Kapetanios, Shin, and Snell (2003).

4.2. TAR/MTAR cointegration test

The long-term relationships between the non-stationary variables can be studied by the cointegration test. TAR/MTAR cointegration tests are used to examine the nonlinear relation between the oil price and stock price indexes.

AlphaShares China Volatility Index is used as threshold in this study. The TAR/MTAR test results are presented in Tables 4. In the cases of Shanghai Composites Index and Shenzhen Composite Index, the threshold values of MTAR model are smaller than the threshold values of TAR model. Most of AIC and SBC of MTAR model are smaller than those in the TAR model. Therefore, MTAR model is better than TAR model in China and Hong Kong stock markets. Table 4 shows that most of the F_c values are significant at 1% or 5% or 10% levels in both TAR and MTAR models. Therefore, the results reject the null hypothesis ($H_0: \rho_1 = \rho_2 = 0$). This represents the existence of cointegration. In the other words, the relationship is stationary in long-term between stock price indexes and spot oil price.

$F_a(H_0: \rho_1 \neq \rho_2)$ in Table 4 is used to test the symmetrical relationships between stock price indexes and oil price in the short-term. F_a of HANG SENG Index /WTI, and F_a of Shenzhen Composite Index /WTI are significant at 5% level in the MTAR models. This indicates that HANG SENG Index and WTI oil price, Shenzhen Composite Index and WTI oil price are symmetrical in the short-term. F_a of Shanghai Composites Index /WTI is not significant both in the TAR and MTAR models. This indicates that there is no symmetrical adjustment between the

residuals of Shanghai Composites Index and WTI oil price.

The results of TAR/MTAR cointegration tests show that the long-term relationships between stock price and oil price are nonlinear for HANG SENG Index, Shanghai Composites Index, and Shenzhen Composite Index, especially HANG SENG Index. Therefore, the nonlinear threshold error correction model can be used to examine the relationships between HANG SENG Index and WTI oil price, between Shanghai Composites Index and WTI oil price, and between Shenzhen Composite Index and WTI oil price.

4.3. Threshold Error Correction Model

The asymmetric threshold error correction model (TECM) with consistent threshold estimates is used to capture the nonlinear relationships between HANG SENG Index and WTI oil price, between Shanghai Composites Index and WTI oil price, and between Shenzhen Composite Index and WTI oil price. Standard T test is used to study whether coefficients of $\Delta STOCK_i$ and ΔWTI are statistically different from zero, and whether the coefficients of the error-correction terms are significantly different from zero. With the appropriate lag lengths being determined by AIC, the results are presented in tables 5 and 6. Table 5 shows the results of dependent variable $\Delta STOCK$ and independent variable ΔWTI . The results of dependent variable ΔWTI and independent variable $\Delta STOCK$ are shown in table 6.

Table 5 shows that the threshold ($\gamma = ASC-VIX$) is 39.79 for HANG SENG Index as dependent variable. Because β_3 and β_4 are not significant, WTI does not influence HANG SENG Index

<Table 4> TAR/MTAR cointegration test between Stock and WTI

		Threshold	F_c	F_a	AIC	SBC
HANG SENG Index / WTI	TAR	0.0000	4.20288**	0.005953	-727.02416	-717.21065
		-0.12240	5.12527***	1.835317	-728.85551	-719.04200
	MTAR	0.0000	4.43083**	0.458040	-727.47705	-717.66354
		0.00326	6.74110***	5.039987**	-732.05557	-722.24206
		Threshold	F_c	F_a	AIC	SBC
Shanghai Composites Index / WTI	TAR	0.0000	1.83160	0.056243	-790.38080	-780.56729
		0.12857	1.90371	0.199942	-790.52477	-780.71126
	MTAR	0.0000	1.83140	0.055855	-790.38041	-780.56690
		0.01179	2.77113*	1.928540	-792.25499	-782.44148
		Threshold	F_c	F_a	AIC	SBC
Shenzhen Composite Index / WTI	TAR	0.0000	2.42472*	0.292425	-411.64612	-401.83261
		0.15263	3.13404**	1.704597	-413.05971	-403.24620
	MTAR	0.0000	2.54744*	0.536732	-411.89082	-402.07731
		-0.01896	4.60609**	4.635305**	-415.98699	-406.17348

1. The critical values of F test for TAR and MTAR are from Enders and Siklos (2001).

2. F_c and F_a denotes the cointegration test and asymmetric test, respectively.

3. ***, **, and * denote the significant levels at 1%, 5%, and 10%.

when ASC-VIX is below the threshold. α_3 is significant at 1% level. Therefore, the lagged one period WTI oil price influences current period HANG SENG Index above the threshold. And the coefficient is negative. A negative coefficient means that WTI oil price and HANG SENG Index will move toward different direction. When HANG SENG Index rises, WTI oil price will fall. The result of short-term causality test and long-term causality test are all significant at 1% level. This indicates that, both short-term and long-term, WTI price is able to influence HANG SENG Index. Both α_5 and β_5 are negative and significant at 1% level. These indicate that WTI oil price and HANG SENG Index can adjust to long-term equilibrium. Table 6 shows that the threshold is 44.34 for HANG SENG Index as independent variable. No matter ASC-VIX is above or below the threshold, the lagged one and two period HANG SENG Index can influence current period WTI oil price. α_1 , α_2 , β_1 and β_2 in Table 6 are all significant. And the coefficients are negative. WTI oil price and HANG SENG Index will move toward different direction. The short-term causality test and long-term causality test are all significant at 1% level. These mean, both in short-term and long-term, HANG SENG Index is able to influence WTI oil price. Another α_5 and β_5 in Table 6 are negative and significant at 1% level; these indicate that WTI oil price and HANG SENG Index can adjust to long-term equilibrium.

Table 5 shows that the threshold is 36.45 for Shanghai Composites Index as dependent variable. α_3 and β_3 both are significant at 10% level. This means no matter ASC-VIX is above or below the threshold; the lagged one-period WTI oil price has influence on the current period Shanghai Composites Index. With negative coefficients means WTI oil price and Shanghai Composites Index move toward different directions. The short-term causality test and long-term causality test are all significant at 1% level. These mean, both in the short-term and long-term, WTI oil price is able to influence Shanghai Composites Index. Both α_5 and β_5 are negative and significant at 1% level. These indicate that WTI oil price and Shanghai Composites Index could adjust to long-term equilibrium. Table 6 shows that the threshold is 45.87 for Shanghai Composites Index as independent variable. No matter ASC-VIX is above or below the threshold, the lagged one period Shanghai Composites Index can influence current period WTI oil price. Both α_1 and β_1 in Table 6 are all significant and the coefficients are negative. WTI oil price and Shanghai Composites Index will move toward different directions. The long-term causality test is significant at 1% level. This means that, Shanghai Composites Index is able to influence WTI oil price in the long-term. Both α_5 and β_5 in Table 6 are negative and significant at 1% level. These indicate that WTI oil price and Shanghai Composites Index can adjust to long-term equilibrium.

Table 5 shows that the threshold is 39.03 for the Shenzhen Composite Index as dependent variable. No matter ASC-VIX is above or below the threshold, the lagged one and two period WTI oil prices have no influence on the current period Shenzhen Composite Index. But the results of short-term cau-

sality test and long-term causality test are all significant at 1% level. These mean that, both in the short-term or long-term, WTI is able to influence Shenzhen Composite Index. Both α_5 and β_5 are negative and significant at 1% level. These indicate that WTI oil price and Shenzhen Composite Index could adjust to long-term equilibrium. Table 6 shows that the threshold is 45.87 for the Shenzhen Composite Index as the independent variable. The lagged two periods Shenzhen Composite Index can influence the current WTI oil price when ASC-VIX is above the threshold. α_2 in Table 6 is significant. But the coefficient is negative. Therefore, WTI oil price and Shenzhen Composite Index will move toward different directions. The long-term causality test is significant at 1% level. This means, in long-term, WTI oil price is able to influence Shenzhen Composite Index. α_5 and β_5 in Table 6 are negative and significant at 1% level. These indicate that WTI oil price and Shenzhen Composite Index can adjust to long-term equilibrium.

The relationship between WTI and HANG SENG Index is more significant than the relationship between WTI and Shanghai Composites Index or Shenzhen Composite Index, when used investor sentiment as the threshold variable. The reasons may be that China has many natural resources and energy, but Hong Kong lacks and totally depends on imported, and therefore HANG SENG index has higher sensitivity to changes in international oil prices; it is to say when investor sentiment variable above threshold value, the degree of oil price impacting on HANG SENG index is deeper than Shanghai Composites Index or Shenzhen Composite Index. The other hand, it may be Hong Kong's economic freedom index is very high, and the account of IPO in Hong Kong Exchanges and Clearing Limited (HKEx) is the most in the world from 2009 to 2011, as a result HANG SENG Index can impact on WTI, and more significant than China.

Another, if only consider Shanghai Composites Index and Shenzhen Composite Index of China, the relationship between Shanghai Composites Index and WTI is more significant than Shenzhen Composite Index under investor sentiment as the threshold variable.

5. Conclusion

The purpose of this research is to study the nonlinear relationship between stock price index and oil price, which is under the different investor sentiment in China and Hong Kong. The variables are daily closing prices, consist of Shanghai Composites Stock Index and Shenzhen Composite Stock Index of China, HANG SENG Index of Hong Kong, Spot Price of WTI crude oil, and used AlphaShares Chinese Volatility Index (ASC-VIX) as investor sentiment variable and as threshold.

The result was the relationship between stock price index and oil price is different under the different investor sentiment, but stock price index and oil price could adjust to long-term equilibrium, and long-term causality test between stock price in-

<Table 5> TECM estimation – Dependent Variable: $\Delta STOCK_i$, Independent Variable: ΔWTI

	HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index
$\alpha 1$ --PLUS	0.190530216*** (3.18848)	0.107666129* (1.75261)	0.042834844 (0.60621)
$\alpha 2$ --PLUS	0.145717514*** (3.46003)	0.144247100*** (3.20667)	0.161415432*** (3.17988)
$\alpha 3$ --PLUS	-0.163442352*** (-6.57152)	-0.034229480* (-1.66154)	-0.025344006 (-1.07514)
$\alpha 4$ --PLUS	-0.035699303 (-1.53190)	-0.021343712 (-1.05713)	-0.008516683 (-0.36266)
$\alpha 5$ --PLUS	-1.264716592*** (-17.02952)	-1.113516662*** (-15.36775)	-0.995461383*** (-12.05529)
$\beta 1$ --MINUS	-0.108300510 (-1.31876)	-0.170258496** (-2.53281)	-0.083628371 (-1.49997)
$\beta 2$ --MINUS	-0.074838963 (-1.38351)	-0.086542680** (-1.92555)	-0.077318280* (-1.90578)
$\beta 3$ --MINUS	-0.049385358 (-1.38744)	-0.058007574* (-1.77460)	-0.045227424 (-1.31674)
$\beta 4$ --MINUS	-0.023948001 (-0.73563)	-0.010654063 (-0.32192)	-0.024095404 (-0.71331)
$\beta 5$ --MINUS	-0.941693603*** (-9.27748)	-0.885078701*** (-10.51471)	-0.889607628*** (-13.13477)
γ	39.79	36.45	39.03
Short-term Causality Test	11.97177***	10.28273***	10.11165***
Long-term Causality Test	15.50827***	6.28096***	4.56332***
AIC	-767.21685	-815.15287	-584.76692
SBC	-708.37188	-756.30790	-525.92195
L-BQ(12)	27.196***	10.893	8.830
L-BQ2(12)	780.094***	343.637***	333.485

1. (): T-Stat

2. ***, **, and * denote the significant levels at 1%, 5%, and 10%

<Table 6> TECM estimation – Dependent Variable: ΔWTI , Independent Variable: $\Delta STOCK_i$

	HANG SENG Index	Shanghai Composites Index	Shenzhen Composite Index
$\alpha 1$ --PLUS	-0.385358137*** (-6.73308)	-0.354982018*** (-5.12961)	-0.386756239*** (-5.58742)
$\alpha 2$ --PLUS	-0.108560627** (-1.97813)	-0.010674238 (-0.15001)	-0.106666221 (-1.49643)
$\alpha 3$ --PLUS	0.243872486*** (4.03611)	0.226807251*** (4.00064)	0.198605278*** (3.51846)
$\alpha 4$ --PLUS	0.191686680*** (4.35017)	0.183614885*** (4.47064)	0.172016884*** (4.21079)
$\alpha 5$ --PLUS	-1.203897327*** (-15.92373)	-1.167427228*** (-16.60088)	-1.156510638*** (-16.63197)
$\beta 1$ --MINUS	-0.275436049*** (-4.15116)	-0.085516984* (-1.74465)	-0.065890160 (-1.59517)
$\beta 2$ --MINUS	-0.172898673** (-2.83118)	-0.028378409 (-0.59186)	-0.021926502 (-0.53640)
$\beta 3$ --MINUS	0.024198074 (0.30630)	0.009406289 (0.13126)	0.010679126 (0.15048)
$\beta 4$ --MINUS	0.036923024 (0.68766)	0.040096273 (0.80996)	0.042073383 (0.85514)
$\beta 5$ --MINUS	-1.037924925*** (-10.72762)	-1.018092807*** (-11.75747)	-1.015853896*** (-11.86213)
γ	44.34	45.87	45.87
Short-term Causality Test	3.91299**	0.02250	2.23930
Long-term Causality Test	11.56455***	6.42951***	6.27499***
AIC	42.35075	-8.33585	-18.26606
SBC	101.19571	50.50911	40.57891
L-BQ(12)	29.135***	22.390**	23.235**
L-BQ2(12)	434.522***	443.190***	489.480***

1. (): T-Stat

2. ***, **, and * denote the significant levels at 1%, 5%, and 10%

dex and oil price were all significant. When used investor sentiment as the threshold variable, the relationship between WTI and HANG SENG Index is more significant than the relationship between WTI and Shanghai Composites Index or Shenzhen Composite Index, and Shenzhen Composite Index is weakest.

The empirical results of the paper offer four major findings. First, there is evidence that the correlation between the stock and oil markets is not constant but time-varying. It tends to increase with the volatility in the market. Second, the daily volatility in each market is very persistent. That is, it varies over time in a predictable manner, conditioned on the past volatilities. Third, there is bidirectional dependence in volatility between the two markets. Innovations that hit either market can affect the volatility in the other market. In other words, conditioned on the persistence and the past volatility in their markets, the past volatility of the stock (oil) market also has predictive power over the future volatility of the oil (stock) market.

The different of relationship between oil price and stock price index under investor sentiment as the threshold variable, may be about one self's resources and energy, but other reasons, example: economic freedom, the account of IPO..., perhaps give deep influences.

References

- Driesprong, G., Jacobsen, B., & Maat, B. (2008). Striking oil: Another puzzle? *Journal of Financial Economics*, 89, 307–327.
- El-Sharif, I., Brown, D., Burton, B., Nixon, B., & Russell, A. (2005). Evidence on the nature and extent of the relationship between oil prices and equity values in the U.K. *Energy Economics*, 27, 819–830.
- Gisser, M., & Goodwin, T.H. (1986). Crude oil and the macroeconomy: tests of some popular notions. *Journal of Money, Credit, and Banking*, 18, 95–103.
- Jones, C., & Kaul, G. (1996). Oil and the stock markets. *Journal of Finance*, 51, 463–491.
- Jones, D.W., Leiby, P.N., & Paik, I.K. (2004). Oil price shocks and the macroeconomy: what has been learned since 1996. *Energy Journal*, 25, 1–32.
- Kanas, Angelos (2012). Modelling the risk–return relation for the S&P 100: The role of VIX. *Economic Modeling*, 29, 795–809.
- Merton, R.C. (1973). An intertemporal capital asset pricing model. *Econometrica*, 41, 867–887.
- Minh Vo. (2011). Oil and stock market volatility: A multivariate stochastic volatility perspective. *Energy Economics*, 33, 956–965.
- Papapetrou, E. (2001). Oil price shocks, stock market, economic activity and employment in Greece. *Energy Economics*, 22, 511–532.
- Park, J., Ronald A., & Ratti, R.A. (2008). Oil price shocks and stock markets in the U.S. and 13 European countries. *Energy Economics*, 30, 2587–2608.
- Rotemberg, J.J., & Woodford, M. (1996). Imperfect competition and the effects of energy price increases on economic activity. *Journal of Money, Credit and Banking*, 28, 549–577.
- Sadorsky, P. (2001). Risk factors in stock returns of Canadian oil and gas companies. *Energy Economics*, 23, 17–28.
- Sari, R., Soytas, U. & Hacihasanoglu, E. (2011). Do global risk perceptions influence world oil prices? *Energy Economics*, 33, 515–524.
- Whaley, R.E. (2000). The investor fear gauge. *Journal of Portfolio Management*, 26, 12–17.